Commentary

Finding \(T_{\text{max}}\) and \(C_{\text{max}}\) in Multicompartmental Models

Yi Rang Han, Ping I. Lee, and K. Sandy Pang

Department of Pharmaceutical Sciences, Leslie Dan Faculty of Pharmacy, University of Toronto, Toronto, Canada

Received May 30, 2018; accepted August 16, 2018

ABSTRACT

Drug absorption data are critical in bioequivalence comparisons, and factors such as the maximum drug concentration (\(C_{\text{max}}\)), time to achieve \(C_{\text{max}}\) (or \(T_{\text{max}}\)), as well as the area under the curve (AUC) are important metrics. It is generally accepted that the AUC is a meaningful estimate of the extent of absorption, and \(T_{\text{max}}\) or \(C_{\text{max}}\) may be used for assessing the rate of absorption. But estimation of the rate of absorption with \(T_{\text{max}}\) or \(C_{\text{max}}\) is not always feasible, as explicit solutions relating \(T_{\text{max}}\) and \(C_{\text{max}}\) to the absorption (\(k_a\)) and elimination rate (\(k\)) constants exist only for the one and not multicompartmental oral model. Therefore, the determination of \(T_{\text{max}}\) or \(C_{\text{max}}\) for multicompartmental models is uncertain. Here, we propose an alternate, numerical approach that uses the point-slope method for the first and second derivative(s) of the concentration-versus-time profiles and the Newton-Raphson iteration method for the determination of \(T_{\text{max}}\) and \(C_{\text{max}}\). We show that the method holds for multicompartmental oral dosing under single or steady-state conditions in the absence of known microconstants, even for flip-flop (\(k_a < \beta\)) models. Simulations showed that the \(C_{\text{max}}\) and \(T_{\text{max}}\) estimates obtained with the Newton-Raphson method were more accurate than those based on the noncompartmental, observation-based method recommended by the US Food and Drug Administration. The %Bias attributable to sampling frequency and assay error were less than those determined by the noncompartmental method, showing that the Newton-Raphson method is viable for the estimation of \(T_{\text{max}}\) and \(C_{\text{max}}\).

Introduction

Drug absorption data are often used in bioequivalence comparisons. The US Food and Drug Administration (FDA) recommends two measures: the area under the curve (AUC) and the maximum drug concentration (\(C_{\text{max}}\)). AUC is the primary and \(C_{\text{max}}\) as the secondary measure. The AUC is known to be independent of \(k_a\), the absorption rate constant under first-order conditions that reflects the extent of drug absorption, and is normally calculated by the trapezoidal rule, which is straightforward for both compartmental and noncompartmental models. The FDA defines \(C_{\text{max}}\) as the maximum drug concentration observed in the sampled blood or plasma data, and is henceforth denoted here as \(C_{\text{max,obs}}\). This \(C_{\text{max,obs}}\) partly serves to provide insight into the rate of absorption.

As previously noted (Endrenyi and Al-Shaikh, 1995; Basson et al., 1996), the \(C_{\text{max}}\) is affected by both the extent and the rate of absorption. In addition, the value of \(C_{\text{max,obs}}\) is further influenced by the frequency of the sampling scheme and the magnitude of the assay errors. As the sampling frequency or assay error is increased, \(C_{\text{max,obs}}\) is increased as well (Endrenyi and Al-Shaikh, 1995). Thus, the comparison between \(C_{\text{max,obs}}\) values that are generated by two divergent sampling schemes or assaying methods may be inappropriate. An alternate measure of the absorption rate is the time to reach \(C_{\text{max,obs}}\) or \(T_{\text{max,obs}}\) which has been suggested to be an unconfounded metric for the rate of absorption (Basson et al., 1996). However, the \(T_{\text{max,obs}}\) is a categorical variable, one that can take on a limited and usually fixed number of possible values and its discriminating power depends strongly on the sampling frequency.

Generally speaking, values for \(C_{\text{max}}\) and \(T_{\text{max}}\) may be obtained by fitting data to compartmental models and then computing those values on the basis of the fitted parameters. In the case of the one-compartment model, the model-guided determination of \(T_{\text{max,1comp}}\) and \(C_{\text{max,1comp}}\) is
clear. The concentration $C_{\text{1comp}}$ may be expressed in terms of the absorption rate constant, $k_a$, and elimination rate constant, $k$, in a biexponential expression, as follows:

$$C_{\text{1comp}} = \frac{k_a F_{\text{sys}} Dose}{V(k_a - k)} \left[ e^{k t} - e^{-k t} \right]$$  (1)

where $V$ is the central volume of the distribution. The condition at $T_{\text{max,1comp}}$ yields the $C_{\text{max,1comp}}$:

$$C_{\text{max,1comp}} = \frac{k_a F_{\text{sys}} Dose}{V(k_a - k)} \left[ e^{k T_{\text{max,1comp}}} - e^{-k t_{\text{max,1comp}}} \right]$$  (1A)

Explicit solution for $T_{\text{max,1comp}}$ is obtained by solving $dC_{\text{1comp}}/dt = 0$ (eq. 2).

$$T_{\text{max,1comp}} = \frac{\ln(k_a/k)}{k_a - k}$$  (2)

The rate constant, $k_a$, may be estimated via curve stripping or with the Wagner-Nelson method (Wagner and Nelson, 1963) to obtain the fraction remaining to be absorbed (FRA), $FRA = \frac{(k_a - k)}{(k_a + k)} = e^{-k t}$. Together with $k$ and $\frac{k_a F_{\text{sys}} Dose}{V(k_a - k)}$, $C_{\text{max,1comp}}$ and $T_{\text{max,1comp}}$ may be calculated (eqs. 1 and 2). The equation holds in the case of flip-flop when $k_a < k$.

The one-compartment analytical equation for finding $T_{\text{max}}$ and $C_{\text{max}}$ can be extended to a scenario with absorption lag time. This may be derived by substituting the observed time ($t + t_{\text{lag}}$) and then solving the analogous problem of the form, $dC_{\text{1comp,lag}}/dt = 0$ in eq. 2. The one-compartment concentration profile is:

$$C_{\text{1comp,lag}} = \frac{k_a F_{\text{sys}} Dose}{V(k_a - k)} \left[ e^{k(t + t_{\text{lag}})} - e^{-k t_{\text{lag}}} \right]$$  (3)

and the associated $T_{\text{max,1comp,lag}}$ is:

$$T_{\text{max,1comp,lag}} = \frac{\ln(k_a/k)}{k_a - k} + t_{\text{lag}}$$  (4)

For the two-compartmental model, a triexponential expression now describes the drug concentration with oral dosing (Gibaldi and Parrier, 1982):

$$C = Ne^{-k t} + Le^{-\alpha t} + Me^{-\beta t}$$

where $N = \frac{k_a F_{\text{sys}} Dose}{V_{1}(\alpha - k) + L} = \frac{k_a F_{\text{sys}} Dose}{V_{1}(\alpha - k) + L}$, $L = \frac{k_a F_{\text{sys}} Dose}{\beta - k}$, and $M = \frac{k_a F_{\text{sys}} Dose}{V_{1}(\alpha - \beta)}$  (5)

An accurate estimation of the parameters is more complex for such a triexponential expression and for the two-compartmental, oral case (eq. 5), several scenarios are possible. First is the scenario when $k_a$ is much faster than the hybrid decay constant $\beta$, i.e., $k_a > > \beta$, and the three exponential components are easily and accurately separated upon curve stripping (Gibaldi and Parrier, 1982). With $k_{12}$, first-order transfer rate from compartment 2 to 1 ($k_{10}$), $k_{10}$, rate constant for removal in compartment 1 and $V_{1}$, the blood/plasma volume of distribution of drug in the central compartment known from intravenous dosing, $k_a$ can be obtained by the Loo-Riegelman method (Loo and Riegelman, 1968) with the usual assumption that elimination occurs in the central compartment. However, for cases when $\beta \geq k_a$, the distinguishing feature of a prominent "hat" or "nose" that occurs with rapid absorption and slow elimination now disappears, and the oral profile shrinks to one resembling that for the one-compartment model (Chan and Gibaldi, 1985). Under this circumstance, the $k_a$ so obtained by the curve stripping procedure is no longer accurate. Even if the parameters for the two-compartment model can be accurately determined, the analysis of $T_{\text{max}}$ and $C_{\text{max}}$ in drugs that exhibit multiple compartmental characteristics, however, cannot be solved explicitly. Often, the Loo-Riegelman method (Loo and Riegelman, 1968) is used first to obtain $k_a$, then $C_{\text{max}}$ and $T_{\text{max}}$, but the microconstants $k_{12}$, $k_{31}$, and $k_{10}$ and the assumption that elimination occurs from the central compartment are required.

In this article, we describe the use of a numerical approach for model-guided determination of $T_{\text{max}}$ and $C_{\text{max}}$ in multicompartmental models, and compare the estimates obtained to those from direct observations ($T_{\text{max,obs}}$ and $C_{\text{max,obs}}$), a method encouraged by regulatory agencies such as the FDA. This method utilizes the second derivative of the concentration-versus-time profile and the point-slope and Newton-Raphson methods. The bias resulting from assay error and sampling frequency, as well as its performance for different $k_a$ values, are discussed.

**Theory and Methods**

**Finding $T_{\text{max}}$ and $C_{\text{max}}$ in Multicompartment Models**

The $C_{\text{max}}$ and $T_{\text{max}}$ for oral dosing for a two-compartment model may be found in a manner similar to that described for the one-compartment. The $C_{\text{max,2comp}}$ occurs when the first derivative (eq. 6) is zero.

$$\frac{dC_{\text{2comp}}}{dt} = N(-k_a)e^{-k t} + L(-\alpha)e^{-\alpha t} + M(-\beta)e^{-\beta t} = 0 \text{ at } T_{\text{max,2comp}}$$  (6)

The above problem may be solved numerically using the Newton-Raphson method (Galantai, 2000), since the second derivative, $\frac{d^2C_{\text{2comp}}}{dt^2}$ may be easily computed:

$$\frac{d^2C_{\text{2comp}}}{dt^2} = Nk_a^2e^{-k t} + L\alpha^2e^{-\alpha t} + M\beta^2e^{-\beta t}$$  (7)

We will illustrate the solution with examples, first with the two-compartment model, then with multicompartmental models, for drugs with a fast versus a slow absorption rate constant.

**Example 1: Oral Dosing for the Two-Compartment Model Drugs**

The following example demonstrates use of this method to obtain $C_{\text{max,2comp}}$ and $T_{\text{max,2comp}}$. The following parameters were selected for the scenario of fast versus slow absorption: $k_a = 2$ (fast) or 0.1 (slow) h$^{-1}$, with common values for $\beta = 0.2$ h$^{-1}$, $\alpha = 0.8$ h$^{-1}$, $k_{12} = 0.5$ h$^{-1}$, and $[F_{\text{sys}}Dose_{\text{sys}}/V_{1}] = 100$. Microconstants $k_{10}$, $k_{12}$, and $k_{31}$ will remain unchanged when $\beta$, $\alpha$, and $k_{12}$ are kept constant, owing to the relations that exist between $\alpha$, $\beta$, $k_{10}$, $k_{12}$, and $k_{31}$ ($\alpha^{\beta} = k_{10}k_{12}$, and $\alpha + \beta = k_{12} + k_{10} + k_{31}$). In fact, the hybrid constants correspond to $k_{10}$, $k_{12}$, and $k_{31}$ values of 0.32, 0.18, and 0.5 h$^{-1}$, respectively. After substitution of the selected values into eq. 5, values of the drug concentration, $C_{\text{comp}}$ in the central compartment, the derivatives $\frac{dC_{\text{comp}}}{dt}$ and $\frac{d^2C_{\text{comp}}}{dt^2}$, and the $T_{\text{max,2comp}}$ estimates are computed and summarized in Table 1. The log-linear concentration-time profile and the two-compartmental model are shown in Fig. 1A showing a prominent "nose," and components of absorption ($k_a$), distribution ($\alpha$), and elimination ($\beta$) for fast absorption, or an apparent one-compartment profile for slow absorption (Fig. 2A). For fast absorption, the $k_a$ may be obtained by the method of residuals. However, for slow absorption (Fig. 2A), $k_a$ may only be estimated with the Loo and Riegelman method (Loo and Riegelman, 1968) when the microconstants $k_{12}$, $k_{31}$, and $k_{10}$ are known. In contrast, for both fast and slow absorption, the Newton-Raphson method (Galantai, 2000) may be used to obtain $T_{\text{max,2comp}}$ without knowledge of the microconstants.

**Fast $k_a$**. A time-expanded view of Fig. 1B showed that the $dC_{\text{2comp}}/dt$ profile intersected at the zero line (obtained when eq. 6 = 0), then continued negatively before rebounding upwards, rising
asymptotically to parallel and approach 0. The first derivative (Fig. 1B), when set = 0 in eq. 6, provided the $T_{max,2comp,true}$ estimate, which equaled 0.97 h graphically from Fig. 1B (see value in Table 1). Such a fast absorption can occur in the case of immediate-release oral dosage forms of drugs (Home et al., 1999; Stuart-Harris et al., 2000), for which the method of point-slope may be applied to provide a $T_{max,2comp,true}$ at convergence. However, if the initial $T_{estimate1}$ was selected from the red zone (Fig. 1C), the numerical method would not converge to $T_{max,2comp,true}$. Hence, it is important to select an appropriate value for $T_{estimate1}$. For good initial estimates between 0 and 1.9 h (where $\frac{dC_{multi}}{dt}$ is 0 in Fig. 1C, with the $T_{max,2comp,true}$ at 0.97 h), the numerical method would converge and yield the true $T_{max,2comp,true}$ value at the minimum. Estimates larger than 1.9 h, however, may lead to the asymptotic value of 0 (Fig. 1C). When all the $T_{estimate}$ values are plotted on the C-versus-time plot (shown as colored points in inset of Fig. 1A), it may be seen that the estimates readily approach $T_{max,2comp,true}$. With known $T_{max,2comp,true}$, the corresponding $C_{max,2comp,true}$ may be calculated (eq. 5).

Slow $k_{a}$. By contrast, the log-linear concentration profile and the two-compartmental model with slow $k_{a}$, shown in Fig. 2A, revealed an apparent, one-compartment characteristic. The absorption constant ($k_{a}$) could be obtained only with the Loo and Riegelman (1968) method with known microconstants $k_{12}, k_{21},$ and $k_{10}$ or simultaneously curve fitting with intravenous data involving these microconstants. Again, the iterative Newton-Raphson method could be readily applied here. By selecting $T_{estimate1}$, finding $a_{1}$ then $T_{estimate2}$, the point-slope procedure was applied to provide $a_{2}$ and $T_{estimate3}$, then eventually $T_{max,2comp,true}$ (Fig. 2C, showing comparable green and red regions as in Fig. 1C), similar to the iteration scheme of Fig. 1 when $k_{a}$ is faster (2 h$^{-1}$). The only difference here was that the procedure provided a “shallow well” when $\frac{dC_{multi}}{dt}$ was set = 0 in eq. 7 (Fig. 2C). The first derivative (Fig. 2C), when set = 0, revealed the true $T_{max,2comp,true}$ estimate of 5.8 h$^{-1}$. The points on $C_{2comp}$-versus-time plot (Fig. 2A) showed the progression in approaching the true $C_{max,2comp}$ and $T_{max,2comp}$ when $k_{a}$ is slow (0.1 h$^{-1}$). Again, when the initial $T_{estimate1}$ (within the shaded green region, Fig. 2C) was appropriately selected before the minimum for the $dC_{2comp}/dt$ curve (12.5 h, or boundary between shaded green and red

<table>
<thead>
<tr>
<th>TABLE 1</th>
</tr>
</thead>
</table>

<table>
<thead>
<tr>
<th>Results (error free; assay error $E = 0$) arising from the Newton-Raphson method for estimating $T_{max}$ from two- or three-compartmental models</th>
</tr>
</thead>
</table>

<table>
<thead>
<tr>
<th>Two-Compartment, Oral ($\alpha = 0.8$ h$^{-1}$, $\beta = 0.2$ h$^{-1}$; $k_{21} = 0.5$ h$^{-1}$; $F_{sys},Dose_{o},V_{1} = 100$)</th>
<th>Slow Absorption ($k_{a} = 0.1$ h$^{-1}$)</th>
</tr>
</thead>
<tbody>
<tr>
<td>$C_{2comp}$</td>
<td>$d_{2Ccomp}/dt$</td>
</tr>
<tr>
<td>$-139e^{-2}$</td>
<td>$-83.3e^{0.8}$</td>
</tr>
<tr>
<td>$d_{2Ccomp}/dt$</td>
<td>$278e^{-2}$</td>
</tr>
<tr>
<td>$T_{max,2comp,true}$ (h)</td>
<td></td>
</tr>
<tr>
<td>$C_{max,2comp,true}$ (µg/mL)</td>
<td></td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Three-Compartment, Oral ($\lambda_{1} = 1.0$ h$^{-1}$; $\lambda_{2} = 0.8$ h$^{-1}$; $\lambda_{3} = 0.2$ h$^{-1}$; $k_{12} = 0.5$ h$^{-1}$; $k_{11} = 1.5$ h$^{-1}$)</th>
<th>Slow Absorption ($k_{a} = 0.1$ h$^{-1}$)</th>
</tr>
</thead>
<tbody>
<tr>
<td>$C_{2comp}$</td>
<td>$d_{2Ccomp}/dt$</td>
</tr>
<tr>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td></td>
</tr>
<tr>
<td>$T_{max,3comp,true}$ (h)</td>
<td></td>
</tr>
<tr>
<td>$C_{max,3comp,true}$ (µg/mL)</td>
<td></td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Multi-compartment, Oral ($\tau = 6$, parameters same as above; $k_{a} = 2$ h$^{-1}$)</th>
<th>Slow Absorption ($k_{a} = 0.1$ h$^{-1}$)</th>
</tr>
</thead>
<tbody>
<tr>
<td>$C_{multi}$</td>
<td>$d_{Cmulti}/dt$</td>
</tr>
<tr>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td></td>
</tr>
<tr>
<td>$T_{max,multi,true}$ (h)</td>
<td></td>
</tr>
<tr>
<td>$C_{max,multi,true}$ (µg/mL)</td>
<td></td>
</tr>
</tbody>
</table>

$Dose_{o}$, oral dose administered; $F_{sys}$, systemic oral bioavailability, product of fraction absorbed ($F_{abs}$), and availabilities of intestine, liver and lung from the first-pass effect; $k_{a}$, first-order transfer rate from compartment 3 to 1 (three-compartment model).
regions in Fig. 2C), then one would obtain the true $T_{\text{max,2comp,true}}$ at convergence. For good initial estimates between 0 and 12.5 h (where $d^2C_{\text{2comp}}/dt^2$ is 0 in Fig. 2C, with the $T_{\text{max,2comp,true}}$ at 5.77 h), the numerical method would converge and yield the true $T_{\text{max,2comp}}$ value at the minimum. Initial estimates larger than 12.5 h, however, may lead to the asymptotic value of 0 (Fig. 2C). Multiple examples of a long $T_{\text{max}}$ have been observed for the slow, oral absorption profiles of drugs such as sertraline (>7 h) (Allard et al., 1999) and zonisamide (5–6 h) (Mimaki, 1998).

Figures 1 and 2 both illustrate that the Newton-Raphson method works well regardless whether $k_a > \beta$, or $k_a < \beta$.

**Example 2: Oral Dosing for Multicompartment Model Drugs**

The general method is applicable not only to examine temporal data after single oral dosing for the two-compartment model cases but also for multiple dosing at steady-state. The concentration for multicompartmental models is shown below, where hybrid constants are denoted as $\lambda_i$: 
\[ C = A_0 e^{-k_a t} + \sum_{i=1}^{m} A_i e^{-\lambda_i t} \tag{8} \]

where

\[ A_0 = \frac{k_F Dose_{po}}{V_1 \prod_{j=1}^{n} (\lambda_j - \lambda_i)} \]

and

\[ A_i = \frac{k_F Dose_{po}}{V_1 (k_a - \lambda_i) \prod_{j=1, j \neq i}^{n} (\lambda_j - \lambda_i)} \]

To demonstrate this, a three-compartment, single oral dose model, with elimination from central compartment, was simulated for the case where \( k_a > \lambda_i \) and \( k_a < \lambda_i \). The selected values are summarized in Table 1, with the method for estimation of \( T_{max,3\text{comp}} \) the same as that for the two-compartment model (see Fig. 3). With known \( T_{max,3\text{comp}} \), the corresponding \( C_{max,3\text{comp}} \) may be calculated.

**Extension of the Newton-Raphson Method for Estimating \( T_{max} \) in Multiple Dosing**

The method may be readily applied to multiple dosing at steady-state, as exemplified below for the two-compartment model, where \( \lambda_i \) is the time elapsed after the last dose here. The domain of the numerical problem is \( 0 \leq u \leq \tau \), where \( \tau \) is the interval of administration.

\[ C_{2\text{comp},SS} = N \left( \frac{1}{1-e^{-k_a \tau}} \right) e^{-k_a u} + L \left( \frac{1}{1-e^{-\alpha \tau}} \right) e^{-\alpha u} + M \left( \frac{1}{1-e^{-\beta \tau}} \right) e^{-\beta u} \tag{9} \]

\[ C_{max,2\text{comp},SS} \] occurs at \( u_{max,2\text{comp},SS} \), the true \( T_{max,2\text{comp}} \) or \( T_{max,2\text{comp},true} \) after the last dose can be computed by solving \( \frac{dC_{2\text{comp},SS}}{du} = 0 \).

\[ \frac{dC_{2\text{comp},SS}}{du} = \left( -\frac{k_F N}{1-e^{-k_a \tau}} \right) e^{-k_a u} + \left( -\frac{\alpha L}{1-e^{-\alpha \tau}} \right) e^{-\alpha u} + \left( -\frac{\beta M}{1-e^{-\beta \tau}} \right) e^{-\beta u} = 0 \tag{10} \]

Here, the value of \( u \), which satisfies the condition for eq. 11 (shown below), occurs at \( u_{max,SS} \), the time elapsed since last dose where maximum concentration \( C_{max,2\text{comp},SS} \) occurs, in the steady state. This method may be extended to multicompartiment multiple dosing systems whose steady-state concentration with dosing interval of \( \tau \) is given by:

\[ C_{SS} = A_0 \left( \frac{1}{1-e^{-k_a \tau}} \right) e^{-k_a u} + \sum_{i=1}^{m} A_i \left( \frac{1}{1-e^{-\lambda_i \tau}} \right) e^{-\lambda_i u} \tag{11} \]

An example of the calculation is shown for the two and three compartments in Fig. 4, and detailed parameter values are summarized in Table 1. The method for estimation of \( T_{max,\text{multi},true} \) is the same as that for all other models described above.
Simulations

Simulations were performed to assess the properties of the proposed method to compare this to noncompartmental analysis, where \(C_{\text{max}}\) and \(T_{\text{max}}\) use observed data as a basis. The pharmacokinetic parameters used in the previous two-compartment model were applied here—\(\alpha\), \(\beta\), and \(k_{21}\) of 0.8, 0.2, and 0.5 h\(^{-1}\), respectively. \([F_{\text{sys}}\text{Dose}_{\text{po}}/V_1]\) was set to 100 [where \(\text{Dose}_{\text{po}}\) is the oral dose administered; \(F_{\text{sys}}\) is systemic oral bioavailability, product of fraction absorbed (\(F_{\text{abs}}\)) and availabilities of intestine, liver, and lung from the first-pass effect]. Blood/Plasma concentration data were simulated using commercial software (MATLAB; The MathWorks Inc., Natick, MA) for the two-compartment, oral model, with elimination in central compartment. The designated hybrid (\(a\) and \(b\)) and \(k_{a}\) and \(k_{21}\) values were used to compute \(N\), \(L\), and \(M\) with eq. 5 and then used to provide values of \(C_{2\text{comp},\text{cal}}\) for each of the designated sampling times. We compared the noncompartmental method via visual identification and our method under different assay errors and sampling frequencies. Simulations were repeated 500 times and accuracy and precision of the two estimated metrics (\(X = C_{\text{max}}, T_{\text{max}}\)), each acquired through two different methods, were evaluated by computing bias and relative S.D.

\[
\text{%Bias} = \frac{\bar{x} - x_{\text{true}}}{x_{\text{true}}} \times 100\% \quad (12)
\]

The relative S.D. (RSD) of the repeat simulations was calculated using eq. 13, where \(x_i\) is the estimated metric for each of the repeats and \(\bar{x}\) is the mean of the estimates and \(N\) is the number of replicate simulations (\(N = 500\)).

\[
\text{RSD} = \sqrt{\frac{\sum_{i=1}^{N} x_i - \bar{x}}{N-1}} \times \frac{1}{\bar{x}} \quad (13)
\]

Assay Error. To understand the effect of assay error, simulations were performed using conditions similar to those proposed by Tothfalusi and Endrenyi (2003). The sampling time was set to a geometric sequence starting at 0.05 and with a geometric ratio of 1.3 for a total of 24 time points. For the two-compartment model with elimination from the central compartment, \(k_{a}\) was set to be 2 h\(^{-1}\). Then, Gaussian noise, with a mean of 0 and relative S.D. of \(\frac{1}{100}C_{2\text{comp},\text{cal}}\) (where \(C_{2\text{comp},\text{cal}}\) is the error-free value and \(E\) is the assay error, 1%–15%) was generated to simulate error-containing concentrations (\(C_{2\text{comp},\text{error}}\)) among the 500 simulations for each of the five designated assay errors (1%, 2%, 5%, 10%, and 15%). The parameters, \(T_{\text{max},2\text{comp}}\) and \(C_{\text{max},2\text{comp}}\) were obtained directly from \(N\), \(L\), and \(M\), \(k_{a}\), \(\alpha\), and \(\beta\), with the set of \(C_{2\text{comp},\text{error}}\) (with error) data from nonlinear fitting of the two-compartment model using MATLAB, which provided a solution for the constants, \(N\), \(L\), \(M\), \(k_{a}\), \(\alpha\), and \(\beta\). These fitted estimates were substituted into eq. 6 and Newton-Raphson numerical method was performed to solve the resulting expression. The noncompartmental \(T_{\text{max},\text{obs}}\) was used as the initial estimate. In cases where proper
convergence was not achieved, the previous time point to $T_{\text{max,obs}}$ was used as the initial estimate and the numerical analysis was repeated. The analysis yielded $T_{\text{max,2comp}}$ and $C_{\text{max,2comp}}$ for comparison with their counterparts in the absence of assay error, $T_{\text{max,2comp,true}}$ and $C_{\text{max,2comp,true}}$. Alternately, the parameters may be used to simulate concentrations at small incremental time intervals to arrive at $C_{\text{max,true}}$ estimates close to the $C_{\text{max,2comp,true}}$ and its corresponding time, $T_{\text{max,2comp,true}}$ were then identified visually.

The %Bias and relative S.D. values of $C_{\text{max,2comp}}$ (Fig. 5, A and C) and $T_{\text{max,2comp}}$ (Fig. 5, B and D) owing to assay error from our proposed model-guided estimation method were much smaller than those from the noncompartmental method, whose basis is merely visual inspection of the observed data for $C_{\text{max,2comp,obs}}$ and $T_{\text{max,2comp,obs}}$. Essentially, there was little bias on $C_{\text{max,2comp}}$ and $T_{\text{max,2comp}}$ for our model-guided estimation method, whereas for the noncompartmental method, the %Bias for $C_{\text{max,2comp,obs}}$ was higher and proportional to the assay error (Fig. 5A). Thus, $C_{\text{max,2comp}}$ values obtained from datasets with different assay errors cannot be compared reliably when the noncompartmental method is used, even though the compartmental model-guided values allow for comparison across different studies.

**Sampling Frequency.** Next, for evaluation of the impact of sampling the frequency, 500 simulations were performed for each for the four sampling frequencies ranging from 5, 10, 20, and 40. The sampling frequency reflects the number of sampling points between 0 and $T_{\text{max,2comp,true}}$ allowing the true $T_{\text{max,2comp}}$ to fall on one of the sampling times. Again, the hypothetical patient data were simulated with the assigned set of $k_a, k_{21}, \alpha$, and $\beta$ values assuming elimination from the central compartment, a specific blood sampling scheme, and a constant 5% assay error. These data were refitted by nonlinear least squares method to obtain the parameters $N, L, M, k_a, \alpha$, and $\beta$, such that the $C_{\text{max,2comp,true}}$ and $T_{\text{max,2comp,true}}$ could be obtained with the Newton-Raphson method, as previously described.

Even as the sampling frequency was increased, $C_{\text{max,2comp}}$ obtained by our method (Fig. 6A) showed no apparent bias, in contrast to the FDA-recommended method, which showed positive bias. Furthermore, the degree of bias increased as sampling frequency was increased, suggesting that $C_{\text{max}}$ cannot easily be compared between datasets that were generated using different sampling frequencies. The positive bias for the visual inspection method is at first counter-intuitive since in the absence of assay errors, $C_{\text{max}}$ should have negative bias as by definition all other concentration values are equal or lower. However, owing to the presence of assay errors, the measured concentrations near the true $C_{\text{max,true}}$ may be greater than $C_{\text{max,true}}$ and since the noncompartmental method whose basis is visual inspection selects the largest concentration, positive bias occurs on average.

In addition, both $T_{\text{max,2comp}}$ and $C_{\text{max,2comp}}$ can be obtained with more certainty (lower relative S.D.) using our method than when obtained directly from the measurements (Fig. 6, C and D). Because population variation was not added to the data, the variation measured here arose from the assay errors and the sampling scheme. Figure 6B shows that the $T_{\text{max,2comp}}$ obtained using the two methods show only minor bias. The small positive bias shown here results from the asymmetric peak shape near $T_{\text{max,true}}$ (i.e., peak is sharper in the left region than the right region). As previously noted (Basson et al., 1996), measuring $T_{\text{max,2comp}}$ is a universal method and the $T_{\text{max,2comp}}$ may be used also under flip-flop kinetics, for which separation of $k_a$ may be difficult without intravenous data. However, the determined $T_{\text{max,2comp}}$ suffers from a higher relative S.D. compared with $C_{\text{max,2comp}}$ (Fig. 6, C and D).

$k_a$. The influence of $k_a$ on the $T_{\text{max,2comp}}$ and $C_{\text{max,2comp}}$ estimates and the applicability of the method were appraised using simulations with five different $k_a$ values: 3, 2, 1, 0.1, and 0.05 h$^{-1}$ (Table 2). The sampling times were made between 0.025 and 48 h (at 0.025, 0.05, 0.1, 0.25, 0.75, 1, 1.5, 2, 2.5, 3, 4, 6, 8, 12, 18, 24, 30, 36, 42 and 48 h) with an assay error of 5% (i.e., assay error $E = 5$). The method was found to be applicable to cases where $k_a > \beta$ and $k_a < \beta$. For all $k_a$ values, the RSD was lower for the two-compartmental fitting-based method, and varied between 2.19% and 5.78% in contrast to the noncompartmental method, whose RSD varied between 3.76% and 23.3% (Table 2).
Discussion

Generally speaking, the model-guided estimation method proposed here using the Newton-Raphson method requires additional analysis in comparison with noncompartmental analysis. The Newton-Raphson numerical method was implemented from scratch and the estimates would converge well when the Tmax,obs was used as initial estimate. Moreover, in various mathematical software applications, zero-finding methods are available as built-in functions (for example, fzero in the case of MATLAB). Not all of these use the Newton-Raphson method, but the functions used are adequate in solving the zero-finding problem numerically (i.e., f(x) = 0 problem). Despite the small amount of added work compared with the noncompartmental analysis method, the model-guided analysis provides reduced bias for the determination of Cmax (Fig. 5A; Fig. 6A; Table 2) and is associated with a higher precision in the determination of both Cmax (Fig. 5C; Fig. 6C) and Tmax (Fig. 5D; Fig. 6D). Moreover, in this analysis, we have not considered the Cmax at fast burst effect of the drug at t = 0, when the burst behaves like an intravenous bolus, owing to the accompanying complexities.

In contrast to the Loo-Riegelman method (Loo and Riegelman, 1968), the model-guided estimation method proposed here does not assume a detailed mechanistic model nor require microconstants, making it suitable even when intravenous data are not available. In addition, our method may be used to analyze data under flip-flop kinetics; this is demonstrated by the example with ka,b (Fig. 2). In fact, our method is applicable generally to all multicompartmental models that are approximated as sum of exponentials under linear conditions. While the number of exponentials is required, explicit knowledge of the detailed compartmental scheme is not required.

One disadvantage of our method is that bias may be introduced when an inappropriate number of exponentials is selected. In cases where the constants are close in values, that is, ka,C2 or ka,C,B, the correct number of compartments will probably be underestimated, and Tmax,true is estimated with increased bias, albeit modest. Under the condition that ka = α or when ka = β, the accurate identification of the Tmax,true is not hampered and may be estimated without bias although we may be underestimating the number of compartments (data not shown). However, when the exponents are sufficiently separated, the appropriate number of exponentials will be easily selected, and the model-guided interpolation will provide a lower bias on the Cmax estimate and a higher precision for Cmax and Tmax. Hence, the method is viable for almost all scenarios. Furthermore, this method may be used even without intravenous data or a priori knowledge about the rank order of the hybrid constants α, β, and ka. Thus, the model-guided estimation method is appropriate for comparative analyses across aggregated datasets available from various sources, with varying collection parameters and limited mechanistic knowledge about the system.

TABLE 2
Effects of varying ka on estimates of Tmax and Cmax from simulated blood data, which contained 5% assay error

<table>
<thead>
<tr>
<th>ka (h⁻¹)</th>
<th>Cmax,true (μg/ml)</th>
<th>Cmax,2comp (μg/ml)</th>
<th>%Bias</th>
<th>RSD</th>
<th>Cmax,true (μg/ml)</th>
<th>Cmax,2comp (μg/ml)</th>
<th>%Bias</th>
<th>RSD</th>
</tr>
</thead>
<tbody>
<tr>
<td>3</td>
<td>70.7</td>
<td>70.9</td>
<td>0.222%</td>
<td>2.54%</td>
<td>72.1</td>
<td>2.04%</td>
<td>3.76%</td>
<td></td>
</tr>
<tr>
<td>2</td>
<td>64.2</td>
<td>64.1</td>
<td>0.077%</td>
<td>2.49%</td>
<td>65.3</td>
<td>1.74%</td>
<td>3.98%</td>
<td></td>
</tr>
<tr>
<td>1</td>
<td>51.9</td>
<td>51.8</td>
<td>0.067%</td>
<td>2.19%</td>
<td>52.8</td>
<td>1.76%</td>
<td>3.90%</td>
<td></td>
</tr>
<tr>
<td>0.1</td>
<td>16.3</td>
<td>16.3</td>
<td>0.338%</td>
<td>2.62%</td>
<td>16.6</td>
<td>2.00%</td>
<td>3.76%</td>
<td></td>
</tr>
<tr>
<td>0.05</td>
<td>10.0</td>
<td>10.1</td>
<td>0.169%</td>
<td>2.69%</td>
<td>10.2</td>
<td>2.09%</td>
<td>3.81%</td>
<td></td>
</tr>
</tbody>
</table>

%Bias was determined using eq. 12.
RSD was determined using eq. 13.
Conclusion

A proper multiexponential fit, combined with the Newton-Raphson method, can provide a better approach for C_{max} and T_{max} estimates than the observational, noncompartmental analysis method recommended by FDA, and is more appropriate when limited data are available.

Authorship Contributions

Participated in research design: Han, Pang.
Conducted experiments: Han.
Contributed new reagents or analytic tools: Han.
Performed data analysis: Han.
Wrote or contributed to the writing of the manuscript: Han, Lee, Pang.

References


Address correspondence to: K. Sandy Pang, Department of Pharmaceutical Sciences, Leslie Dan Faculty of Pharmacy, University of Toronto, 144 College Street, Toronto, Ontario M5S 3M2, Canada. E-mail: ks.pang@utoronto.ca